

2.6 Inverse chi-squared statistic

Our starting distribution is the Gaussian

$$\text{prob}(x) = \frac{1}{\sqrt{2\pi}\sigma} \exp\left[-\frac{(x-\mu)^2}{2\sigma^2}\right]$$

and so for n data X_i we have the posterior distribution for μ and σ

$$\text{prob}(\mu, \sigma | \text{data}) \propto \left(\frac{1}{\sqrt{2\pi}\sigma}\right)^n \exp\left[-\frac{\sum_{i=1}^n (X_i - \mu)^2}{2\sigma^2}\right] \text{prob}(\mu, \sigma).$$

There are reasonably strong theoretical arguments for taking the combined prior for the location and scale parameters to be the Jeffreys prior (see Lee, 1997)

$$\text{prob}(\mu, \sigma) \propto \frac{1}{\sigma}$$

with obvious caveats about the divergent nature of the distribution.

Given this, if we wish to know the posterior distribution of the mean without regard to the possible values of σ , we may simply marginalize it out:

$$\text{prob}(\mu | \text{data}) = \int \text{prob}(\mu, \sigma | \text{data}) d\sigma.$$

A neat trick to tidy up the expression is to use

$$\sum_{i=1}^n (X_i - \mu)^2 = S + n(\langle X \rangle - \mu)^2$$

where

$$S = \sum_{i=1}^n (X_i - \langle X \rangle)^2$$

and the angle brackets denote the simple arithmetic mean of the data.

The marginalization is now a standard integration (change variables to $u = 1/\sigma$) and we find that

$$\text{prob}(\mu | \text{data}) \propto \frac{1}{(S + n(\langle X \rangle - \mu)^2)^{n/2}}$$

which shows clearly how the most probable value of the mean concentrates near the arithmetic mean; the scatter in the data, which must diminish this, appears in the number S – and the parameter σ has of course vanished. This is very close to a standard t-distribution, which as we will see in Chapter 4, turns up naturally when we examine differences of means.

By the same methods we can eliminate the parameter μ , marginalizing it out to obtain

$$\text{prob}(\sigma \mid \text{data}) \propto \frac{1}{\sigma^{n+1}} \exp \left[-\frac{S}{2\sigma^2} \right].$$

This has the form of an “inverse chi-square distribution”, meaning that changing variables to $u = 1/\sigma$ will give a standard chi-square distribution. As in the previous case, we see that the “nuisance parameter” (μ this time) has conveniently vanished, its effect being mediated through the number S again.

In fact the two important numbers S and $\langle X \rangle$ tell us everything we need to know about the data. They are called sufficient statistics, and we will meet them several times in the next Chapter for this reason.